Methods for Achieving Equivalence of Samples in Cross-National Surveys: The European Social Survey Experience

Peter Lynn¹, Sabine Häder², Siegfried Gabler², and Seppo Laaksonen³

Most surveys carried out at national or subnational level involve a single sample design and sampling frame. Where multiple frames are used, they are typically used to access different subpopulations. For cross-national surveys, however, it is usually necessary to use a different design in each nation to sample from an analogous national population. Cross-national sampling frames are rare. In this article, we describe procedures used to obtain equivalence of sample designs in 22 nations in Round 1 of the European Social Survey (ESS). We evaluate the implementation of the procedures and we summarise lessons for the design of cross-national surveys. We focus particularly on novel aspects of the procedures. These include specification of national sample sizes in terms of "effective sample size" and provision of guidelines on how to predict design effect components and how to use the predictions to determine the necessary sample size. We also discuss procedures for interaction between the various parties involved: the ESS central co-ordinating team, the ESS sampling panel, national co-ordinators and field work organisations.

Key words: Design effects; effective sample size; intra-cluster correlation; sample design; sampling frames.

1. Introduction

Compared with surveys carried out within a single nation, cross-national surveys involve an extra layer of complexity in terms of both organisation and design (Lynn et al. 2006). Special procedures are required in order to derive and implement appropriate standards for design and implementation (Lynn 2003). This is particularly important in the case of sample design, as available sampling frames and other constraints tend to vary between nations (see e.g., Adams and Wu 2002; Cornelius 1985; Le 1993). Additionally, field work is often organised at the national level, resulting in the involvement of many persons and organisations in the survey process. It is important to find ways to ensure the comparability of the sample designs used in the different nations, as substantive comparisons between nations, or between groups of nations, are often a key objective of the survey.

In this article, we first discuss the objectives of sample design for cross-national surveys (Section 2). We then describe the principles and requirements for sample design that were developed for the European Social Survey (ESS) in order to meet these objectives (Section 3).

¹ Institute for Social and Economic Research, University of Essex, Wivenhoe Park, Colchester, Essex CO4 3SQ, UK. Email: plynn@essex.ac.uk

² Zentrum für Umfragen, Methoden und Analysen (ZUMA), Postfach 12 21 55, 68072 Mannheim, Germany. Email: sabine_haeder@zuma-mannheim.de; gabler@zuma-mannheim.de

³ PO Box 68, FIN-00014 University of Helsinki, Finland. Email: seppo.laaksonen@helsinki.fi

In particular, these include a requirement to predict design effects and to use these predictions in determining national sample sizes. Given the large number of persons and organisations typically involved in cross-national surveys, effective procedures for implementing the agreed principles and requirements are of paramount importance. The procedures used on the ESS are described in Section 4, and some of the strengths and weaknesses are pointed out. In Section 5, we evaluate the requirements regarding design effects due to both unequal selection probabilities and clustering and make some suggestions as to how such requirements might be better stated in the future. Section 6 concludes.

2. Sample Design for Cross-National Surveys

To enable comparisons between nations, we suggest that national sample designs for cross-national surveys must meet two fundamental criteria:

- The study population must be equivalent in each nation. In practice, this will usually
 mean that the same population definition is applied in each nation and that no or only
 minimal under-coverage can be permitted;
- Sample-based estimates must have known and appropriate precision in each nation. In practice, "known" precision means that a strict probability sample design must be used, and those aspects of sample design that affect precision (selection probabilities, stratum membership, PSU membership) must be available on the microdata to permit estimation of standard errors; "appropriate" precision may mean a) meeting some minimum precision requirement in order for the estimates to be useful, and b) aiming for similar precision in each nation, as this would represent an effective allocation of resources if a prime objective was to make between-nation comparisons.

To best meet these criteria, it is likely that details of the sample design will vary between nations (Le and Verma 1997). The goal is functional equivalence, not replication of parameters of the sample design. As Kish (1994, p. 173) wrote, "Sample designs may be chosen flexibly and there is no need for similarity of sample designs. Flexibility of choice is particularly advisable for multinational comparisons, because the sampling resources differ greatly between countries. All this flexibility assumes probability selection methods: known probabilities of selection for all population elements." Therefore, an optimal sample design for a cross-national survey should consist of the best probability sample design possible in each nation, where "best" can be interpreted as an optimum trade-off between cost and precision. The choice of a specific national design depends on the available frames, experiences, other constraints such as those that may be imposed by the national legal infrastructure and, of course, costs of sample selection and data collection (Häder and Gabler 2003). If adequate estimators are chosen, the resulting estimates can be compared using appropriate statistical tests.

3. Requirements of Sample Design for the European Social Survey

3.1. The European Social Survey

The ESS is an academically-driven social survey designed to chart and explain the attitudes, beliefs and behaviour patterns of Europe's diverse populations. In parallel with its substantive

aims, it aims also to provide a model of best practice in methodology and to contribute to the improvement of methodological standards. Further details of the background and objectives of the survey can be found at www.europeansocialsurvey.org. The ESS is funded via the European Commission's 5th and 6th Framework Programmes, with supplementary funds from the European Science Foundation. In each participating nation, the cost of data collection and the appointment of a national co-ordinator (NC) is funded by the national research council or equivalent body. An important principle of the survey is that the data are made freely available as soon as possible: no one involved in the survey has advance access and there are no restrictions on access. Data can be downloaded from http://ess.nsd.uib.no.

The ESS consists of regular "rounds" of data collection, with each round involving an independent cross-sectional sample in each nation. Survey rounds take place at 2-year intervals. The first round of field work took place in September-December 2002 (in a few nations fieldwork was not completed until 2003.). There is a core questionnaire (approximately 30 minutes) that is administered in every round, along with modules of questions (in Round 1, two modules of approximately 15 minutes each) that will change from round to round. (In Round 1, mean interview length ranged from 51.8 minutes in both Spain and Italy to 69.7 minutes in Sweden.) Nations are not asked to commit themselves to more than one round at a time, though of course continued participation is actively encouraged. Twenty-two nations participated in Round 1, namely the fifteen member states of the EU plus four accession states that became member states subsequently in 2004 (Czech Republic, Hungary, Poland, Slovenia) and three non-EU members (Israel, Norway, Switzerland). All interviews are carried out face-to-face.

The ESS methodological functions have been organised into ten work-packages, overseen by a central co-ordinating team (CCT). One of the work-packages deals with sampling. This work-package is being carried out by a panel of four sampling experts which, at least for the first two rounds of the survey, consists of the four authors of this article (from December 2001 until June 2002, Susan Purdon was also a member of the panel). The panel developed the requirements for participating nations – which will be described in the remainder of this section under five broad headings – and then co-operated with participating nations in developing acceptable sample designs.

3.2. Population Definition and Coverage

The target population for each participating nation is defined as all persons 15 years or older resident in private households within the borders of the nation, regardless of nationality, citizenship, language or legal status. (In countries in which any minority language is spoken as a first language by 5% or more of the population, the questionnaire must be translated into that language.) It is worth noting in passing that this definition was subject to considerable discussion by a 21-country steering group (European Science Foundation 1999) prior to agreement. Concerns raised related almost exclusively to cases in which implementation of the definition would represent a departure from normal practice in a particular country or in which practical difficulties were caused by the nature of the available sampling frames — not to any conceptual or substantive issues.

Thus, the requirement for sample design was that every person with the defined characteristics should have a nonzero chance of selection. In practice, the quality of available frames – e.g., coverage, updating and access – differs between nations, so careful evaluation of frames was necessary to assess the likely extent of under-coverage and ensure that any coverage bias was likely to be minimal.

Among others, we found the following kinds of frames:

- nations with reliable lists of *residents* that are available for social research such as the Danish Central Person Register that has approximately 99.9% coverage of persons resident in Denmark;
- nations with reliable lists of households that are available for social research such as the "SIPO" database in the Czech Republic, that is estimated to cover 98% of households;
- nations with reliable lists of addresses that are available for social research such as
 the postal delivery points from "PTT-afgiftenpuntenbestand" in the Netherlands and
 from the "Postcode Address File" in the UK;
- nations without reliable and/or available lists such as Portugal and Greece.

Drawing a sample is most complicated if no lists are available. In this instance area sample designs (Särndal et al. 1992) were usually applied, in which the selection of a probability sample of small geographical areas (e.g., Census enumeration areas within municipalities) preceded a complete field enumeration of households or dwellings within the sampled areas, from which a sample was selected. In nations where this approach was used (e.g., Greece), the sampling panel insisted that the selection stage should be separated from the enumeration and carried out by office staff or supervisors who had not been present for the enumeration. An alternative to area sampling in this situation is the application of random route sampling, about which some survey organisations were enthusiastic. The basic idea of random route sampling is that within each sampled PSU one address is selected by a random method to serve as a starting point and the interviewer then follows rules that specify the route he or she should take from there, sampling systematically using a prespecified interval (Häder and Gabler 2003). The question here, however, is the extent to which random routes can be judged to be "strictly random." That depends on both the rules for the random walk and the control of the interviewers by the fieldwork organisation in order to minimise interviewer influence on selection. A rigorous version of random route sampling was permitted in one country (Austria).

Even in countries where reliable lists exist, some problems had to be solved. For example, in Italy there is an electoral register available. But it contains, of course, only persons 18 years or older (and only those who are eligible to vote). Therefore, it had to be used as a frame of addresses. This had not been attempted before and there were practical problems to be overcome, not least the fact that persons at the same address do not necessarily appear together on the list, making it difficult to ascertain the selection probabilities of addresses. A new method was developed, resulting in known nonzero selection probabilities for all target population members living at addresses where at least one registered elector resided (Lynn and Pisati 2007). Thus, under-coverage, while not zero, was restricted to persons at addresses with no registered electors. In Ireland, similar issues arose in the use of the electoral registers as a frame. In countries with population

registers, people with illegal status will be excluded because they are not registered. The practical task for the sampling panel was to ensure that levels of under-coverage were kept to an absolute minimum by considering all possible frames and evaluating the properties of each with respect to the ESS population definition. Additionally, all departures from the ideal were documented carefully and are available on the ESS website.

3.3. Response Rates

Nonresponse is the next problem with regard to achievement of the objective to represent the target population. A carefully drawn sample from a perfect frame can be devalued if nonresponse leads to bias. Therefore, it is essential to plan and implement adequate field work strategies to minimise noncontacts and refusals. For the ESS a target response rate of 70% was fixed. This would be particularly challenging for some countries where response rates between 40 and 55 per cent are common (Lyberg 2000). Nevertheless, it was felt that a realistic but challenging target should encourage maximum efforts. Additionally, the ESS required that noncontacts should not exceed 3% of eligible sample units (addresses, households or persons, depending on the sampling frame), that at least four personal visits must be made to a sample unit before noncontact was accepted as an outcome, and that the field period must last at least 30 days. Definitions of outcome categories and response rate were also supplied.

While these targets and constraints were felt to be necessary, they were not expected to be sufficient to ensure high response rates, let alone similar (and small) nonresponse bias in each nation, which was the ultimate goal. The sampling panel and the CCT offered extensive advice on techniques for increasing response rates such as advance letters, toll-free telephone numbers for sample members, calling patterns, incentives, training of interviewers in response-maximisation techniques and in doorstep interactions, etc.

3.4. Sample Selection Methods

We have already argued that strict probability sampling is a necessary prerequisite for cross-national comparability. Without it, the second of the two fundamental criteria introduced in Section 2 cannot be met. However, partly as a measure to overcome the fear of nonresponse bias, many survey organisations habitually implement replacement of noncooperative or unreachable primary sampling units, households or target persons by others. There are many varieties of replacement (Vehovar 2003; Lynn 2004; Rubin and Zanutto 2002), but none of them meet the requirement for probability sampling. Another important disadvantage of replacement in the field is that it can reduce the effort made by interviewers to gain a response at the original addresses/households (Chapman 1983; Elliot 1993).

For the ESS, replacement of nonresponding households or individuals (whether "refusals" or "noncontacts") was not permitted in any circumstances. However, in exceptional circumstances replacement was permitted at the first stage of sampling. Administrative considerations may mean that addresses cannot be obtained for specific sampled areas (e.g., a particular municipality may refuse to grant access to the list, or be unable to co-operate within the available time). In these exceptional cases, provided the circumstance applied to only a very small proportion of the sampled PSUs, replacement with PSUs randomly selected from the same strata was allowed. (This happened in only one nation.)

3.5. Effective Sample Size

The ESS requirement was for a minimum estimated effective sample size of 1,500 completed interviews and a minimum of 2,000 actual interviews. (An exception was made for nations with a total population of less than 2 million, recognising that resources for funding surveys are considerably constrained in such nations. For such nations, the minimum requirement was an effective sample size of 800 and an actual sample size of 1,000. In practice, this applied to only 2 of the 22 nations participating in Round 1: Slovenia and Luxembourg.) Explanation was provided as to what was meant by effective sample size and how it should be predicted. This involved predicting, under certain simplifying assumptions, the design effect due to unequal selection probabilities ($DEFF_p$) and the design effect due to clustering ($DEFF_C$). Additionally, realistic estimates of response rate and eligibility rate were required in order to calculate the sample size to select in order to produce the target number of completed interviews.

This seems a reasonable approach to sample size determination given that it should be possible to predict the determinants of design effects within reasonable bounds for a survey like the ESS. The aspects of the survey that make this possible are 1) relatively low – and relatively stable over time – expected correlations between survey variables and PSUs (because it is a household survey and PSUs are defined by geography); 2) relatively small variation in selection probabilities (unlike, say, an establishment survey with probabilities proportional to a size measure); 3) prior estimates in several countries for similar variables on surveys with similar designs. Additionally, a repeating survey like ESS offers the opportunity to revise predictions at each round based on estimates from previous rounds. There might be a case for a different approach – perhaps with a focus on avoiding worst-case scenarios – in a very different survey situation where design effects might be far less stable.

For most of the National Co-ordinators (NCs) and survey organisations, the concepts of effective sample size and design effect were completely new. The sampling panel invested considerable time and effort in explaining the concepts and helping NCs to estimate design effects. This often involved the NCs seeking out statistical information with which they had no familiarity, such as the national distribution of the number of persons aged 15 or over in a household. In return, we encountered almost universal enthusiasm when it came to understanding the concepts and meeting the requirements. Feedback obtained by the CCT from the NCs after Round 1 suggested that the sample design process had been perceived as one of the most productive and useful aspects of taking part in the ESS and that several nations felt that both knowledge and methods had been improved in their country. Because this process was, to our knowledge, novel and appeared to have a considerable effect, we will describe the requirements and how they were implemented in some detail.

3.5.1. Design Effect Due to Unequal Selection Probabilities ($DEFF_p$)

The ESS guidelines suggested that $DEFF_p$ should be predicted as follows:

$$D\tilde{E}FF_P = \frac{m\sum_{i=1}^{I} m_i(w_i^2)}{\left(\sum_{i=1}^{I} m_i w_i\right)^2} \tag{1}$$

where m_i and w_i denote respectively the number of interviews and the design weight associated with the i^{th} weighting class. (This can be expressed equivalently as $1 + cv_w^2$, where cv_w is the coefficient of variation of the weights – see e.g., Kish 1992.) This is a simplification (Gabler et al. 1999) that assumes $Var(y_{ij}) = \sigma^2$, where $Var(y_{ij})$ is the variance of the target variable y over respondents j within weighting class i. (Kish (1995) refers to weights for which this assumption holds as "random weights.") This may be a reasonable assumption in many cases, but we would note that the extent of departure from this assumption will depend partly on the source of variation in design weights, which in turn will differ between nations.

In some nations, it is necessary to select the sample in stages, with the penultimate stage being addresses or households. In this case, each person's selection probability depends on the household size (number of persons aged 15 or over). The guidelines illustrated estimation of (1) with a hypothetical example of an address-based design of this sort, where the weighting classes were defined by the possible values of the number of persons aged 15 or over resident at an address. A fair number of nations used such an address-based design: Czech Republic, Greece, Ireland, Israel, The Netherlands, Portugal, Spain, Switzerland, and UK. There was considerable variation between these nations in $D\tilde{E}FF_P$ due to differences in the estimated household size distribution.

Another reason for unequal selection probabilities is that minority groups are over-sampled for substantive reasons. Examples of this are Germany, where the East German population is over-sampled, and Israel, where the Arab population is over-sampled. A third reason is that certain strata (typically, the largest cities) may be over-sampled in anticipation of lower response rates, though in principle this should not affect the variance of estimates as it will lead to equal inclusion probabilities if the response rate predictions turn out to be accurate.

A fourth source of variation in selection probabilities occurs in countries where the PSUs are selected with probability proportional to a proxy size measure which does not correlate perfectly with the units sampled at the subsequent stage. Examples include Israel and Ireland, where the PSU size measures were numbers of persons and the selected units were households and addresses respectively.

It should be noted that $D\tilde{E}FF_P$ may vary over sample subgroups. The ESS guidelines are concerned only with the precision of estimation for total sample estimates, but in practice much analysis involves subgroup estimation. Consequently, variation between key subgroups in $D\hat{E}FF_P$ warrants investigation.

3.5.2. Design Effect Due to Clustering ($DEFF_C$)

The cluster sample size and the intra-class correlation also influence the design effect. Following Kish (1987), the ESS guidelines suggested that $DEFF_C$ should be predicted as follows:

$$D\tilde{E}FF_C = 1 + (\bar{b} - 1)\rho \tag{2}$$

where \bar{b} is the mean number of interviews per cluster and ρ is the intra-cluster correlation. (See Lynn and Gabler 2005 for discussion of alternatives to the use of \bar{b} in the prediction of $DEFF_{C}$.) Expression (2) implies that, if cost were not a consideration, the cluster sample size should be chosen as small as possible. The larger the average cluster size, the more

interviews have to be conducted to reach the minimum effective sample size of 1,500. The challenge, therefore, is to find the combination of \bar{b} and n that delivers the desired effective sample size for the lowest cost. Participating nations were encouraged to seek estimates of ρ from other surveys in their country if possible or alternatively to assume $\rho=0.02$ (a value that was chosen as an approximate mean of estimates for attitude measures from surveys in a small number of countries). In practice, ρ will take different values for different statistics and can also vary between subgroups for any particular statistic, but the ESS sample design requirements were stated only in terms of the total sample and only in terms of a "typical" ρ .

Considerable variation in $D\tilde{E}FF_C$ was observed, primarily because of the variation in proposed cluster sample size.

3.5.3. Combined Design Effect

The ESS guidelines suggested that the total design effect should be predicted as:

$$D\tilde{E}FF = D\tilde{E}FF_P \times D\tilde{E}FF_C \tag{3}$$

This of course ignores any design effect due to stratification of the sampling frame, but as this is generally modest in magnitude and beneficial in direction (i.e., less than one), ignoring this effect was felt to both simplify the calculation and insert a little conservatism into the required sampled size. Expression (3) also assumes no association between the weights and the clusters – see Lynn and Gabler (2005). Predictions of total design effect varied greatly between nations. At one extreme, there were three nations where both component design effects, and hence the overall design effect, were 1.00 (Denmark, Finland, Sweden). There were three nations with clustering but no variation in selection probabilities (Belgium, Hungary, Slovenia) and two with variation in probabilities but no clustering (Luxembourg, The Netherlands); and for the other 14 nations, both component design effects were predicted to be larger than one, in most cases substantially larger. The largest predicted design effects were 1.59 for Israel, 1.58 for France, 1.55 for the UK and 1.52 for Germany. Variation in predicted design effects can be seen to be primarily influenced by two factors: the nature of available sampling frames (e.g., population registers that permit equal-probability sampling versus address-based methods) and the interaction between geographical spread of the population, field costs and available budget (which largely determine the extent of clustering, if any).

3.6. Documentation

Comprehensive and clear documentation of all relevant methodological aspects of the survey was demanded. At the level of sampling units, this meant that indicators of sampling stratum, primary sampling unit and the selection probability at each stage of sampling should be included on a micro-level data file that carried the same identifiers as the questionnaire and other data files. A detailed file specification was provided. Supply of these data would allow the application of design weights and the use of appropriate methods for the analysis of data from a complex survey.

Metadata to be supplied and made freely available included a detailed description of the sample design and sample selection process and a clear account of any ways in which the design fell short of the ideal, for example where the frame may have suffered some undercoverage.

4. Developing and Implementing Sample Designs

Providing clear written requirements for participating nations is necessary but does not guarantee that they will be met. In this section we describe the ESS process for developing acceptable sample designs in each nation. The process has two features which we think are particularly important. The first is that it is co-operative rather than authoritarian. The sampling panel saw its prime role as the provision of advice and assistance where needed. The role of monitoring the design and implementation was kept as implicit as possible. The second important feature is that interaction was intensive. Regular contact was made between relevant parties and this promoted the spirit of co-operation and enabled the building of rapport. To our knowledge, this process is unique amongst cross-national social surveys.

As already mentioned, the functions relating to sample design (development, agreement, documentation, and evaluation) are the responsibility of a small panel. A deliberate decision was made to keep the panel small in order to minimise formality and engender the spirit of a team of co-researchers rather than a committee. At the first meeting of the panel, in December 2001, an allocation of Round 1 participating nations to panel members was agreed, so that each member would liaise with and support five or six nations. Panel members contacted "their" National Co-ordinators (NCs) asking for information about the foreseen sampling design. Then, a process of co-operation began. In most cases, the sampling panel member worked closely with the NC; in some cases, the NC preferred instead that the panel member should work directly with the survey organisation while keeping the NC informed. Most of the communication with NCs and survey organisations was done by email, but special visits by panel members to participating nations were allowed where this was thought likely to be particularly beneficial (five such visits were made). Additional face-to-face contact took advantage of other occasions such as conferences and the regular meetings of all NCs with the Central Co-ordinating Team (CCT).

In many countries completely new designs had to be developed to meet the ESS requirements. In other countries, it was only a matter of clarifying details. In particular, support in calculating the effective sample sizes was often necessary. The amount of time and effort committed by the sampling panel members therefore varied greatly over the nations.

Once all questions were clarified and the panel member was satisfied with the proposed design of a country, it was deemed ready for "signing off." The panel had developed a sample form where details of the design of each country had to be completed in a systematic way. This task was done by the panel member for "his/her" countries, thus ensuring the use of standardised terms and ensuring that the design was clearly defined (otherwise the panel member was not able to describe it in the form). Then the panel member presented the form to the other panelists. If all of them agreed, the design was "signed off." Thus, the decision to sign off a design was always made by the whole team together. Otherwise, discussion with the NC had to carry on. This happened in several cases where the panel did not initially agree with the design or requested additional information. Once signed off, the sample form was circulated to the CCT and eventually

incorporated into the sampling panel's report on Round 1. During the period of the development of designs for Round 1 (December 2001 to October 2002), the panel met on three occasions – in London, Mannheim, and Helsinki – to discuss general principles (for example whether random route methods could ever be acceptable – and if so, in what circumstances), processes, and specific designs.

5. Evaluation of the ESS Procedures

5.1. Process

The process described in Section 4 above was felt by all parties to have been a great success. Though the final designs do not strictly meet the stated requirements in every respect in all cases, no unanticipated major departures have yet been identified. Additionally, all known departures are documented and this information is in the public domain on the ESS website. In some countries, the implemented design represented a significant breakthrough in survey standards (see Section 5.8 below). This would not have been possible without such an intensive process of co-operation.

5.2. Predictions of $DEFF_P$

The requirement to predict $DEFF_P$ was novel for all but a few of the participating nations. In several cases, this made the prediction subject to quite heroic assumptions because the necessary information was not available. Uncertainty about the distribution of selection probabilities was of three sorts:

- a Uncertainty about the distribution of household size (number of persons aged 15 or over);
- b Uncertainty about the relationship between a proxy size measure used at the PSU level and the actual size measure of relevance;
- c Uncertainty about the relationship between a proxy size measure used at the household/address level and the actual number of persons aged 15 or over.

There is of course a fourth source of variation in selection probabilities, namely planned over-sampling of domains. But in no case was there uncertainty about the ratios of selection probabilities involved. Over-sampling of this kind was performed in four countries to combat expected response rate differences (The Netherlands, Poland, Portugal, Spain) and in three countries for substantive reasons (Israel, UK, Germany). In all seven cases, the domains were defined by geographical areas. If response rate predictions are accurate, only in the latter three cases will the variance of estimates be affected.

In Table 1 we present the prefieldwork prediction of $DEFF_P$ ($D\tilde{E}FF_P$) and the post-fieldwork estimate of $DEFF_P$ ($D\hat{E}FF_P$) calculated using Expression (1), for each nation that was subject to at least one of the three sources of uncertainty described above. There are 13 such nations; six nations used equal-probability designs and the remaining three (Germany, Norway, and Poland) had complete control over the variation in selection probabilities (all used population registers as the frame), except in so far as nonresponse may have caused the distribution amongst the responding sample to differ from the distribution amongst the selected sample – a source of variation that we shall ignore here.

Table 1.	Predicted	and	Estimated	Values	of DEFF _F

		Source of uncertainty (see text of Section 5.2)			$D ilde{E}FF_P$	$D\hat{E}FF_{P}$	$D ilde{E}FF_P/D\hat{E}FF_P$
		a) Household size distribution	<i>b</i>) Proxy PSU size	c) Proxy household size			
Austria	AT	X			1.25	1.25	1.00
Switzerland	CH	X			1.25	1.21	1.03
Czech Republic	CZ	X			1.16	1.25	0.93
Spain	ES	X			1.16	1.22	0.95
France	FR	X			1.30	1.23	1.06
Greece	GR	X	X		1.18	1.22	0.97
Ireland	ΙE	X			1.03	1.04	0.99
Israel	IL	X	X		1.30	1.56	0.83
Italy	IT		X	X	1.01	1.16	0.87
Luxembourg	LU			X	1.40	1.26	1.11
The Netherlands	NL	X			1.19	1.19	1.00
Portugal	PT	X			1.10	1.83	0.60
United Kingdom	UK	X			1.23	1.22	1.01

The first three columns of the table indicate which sources of uncertainty applied in which nation. With use of Expression (1) for both $D\tilde{E}FF_P$ and $D\hat{E}FF_P$, any difference in the calculated values is due solely to differences between the predicted and observed distribution of design weights.

We would note that type b) uncertainty is unlikely to have a large effect, as the proxy measures used in these three cases are likely to be highly correlated at the PSU level. For example, in Italy municipalities were selected with probability proportional to the number of residents aged 18 or over according to the municipal population registers ("elenco civile"), whereas for that component of the design to be self-weighting, the size measure should have been the number of addresses at which at least one person is a registered elector.

In most cases, then, the uncertainty concerned the distribution of number of persons aged 15 or over at a household/address (source a). There were three nations where this was predicted from recent previous surveys that had used a similar selection method, though in all three cases a different age cut-off had been used (either 16 or 18): Switzerland, Netherlands and UK. In these three cases, differences between $D\tilde{E}FF_P$ and $D\hat{E}FF_P$ are small (the slightly larger difference for Switzerland may possibly be due to variability introduced by a low response rate: Switzerland had the lowest response rate of all nations in Round 1 of the ESS: just 33%, compared to a mean response rate of 63% in the other 21 nations). In the other eight nations that used address-based sampling, no recent survey had used a similar method, so greater guesswork was involved, often involving extrapolations from population estimates or census estimates. In some cases these provided only an estimate of mean household size, in some cases they used a different lower age limit and in some cases they were several years out of date. This resulted in considerable variability in the accuracy of $D\tilde{E}FF_P$ as a prediction of $D\hat{E}FF_P$. For these eight nations, the predictions were too low, with the exception of Austria and France. In the case of France, the variation in household size turned out to be less than predicted, a difference which could have been partly caused by differential nonresponse.

5.3. Predictions of \bar{b}

The mean number of sample units per PSU in the selected sample was of course determined by the design. However, prediction of \bar{b} , the mean number of respondents per PSU, relied upon prediction of the response rate. In practice, \bar{b} differed from the predicted value due to response rates that differed from expectations in several countries. In most cases, the achieved response rates were lower than the predictions, but in some they were higher. The greatest proportionate under-prediction was in Greece ($\bar{b} = 4.8$; $\bar{b} = 5.9$), while the greatest over-prediction was in Italy ($\bar{b} = 18.0$; $\bar{b} = 11.0$), followed by France ($\bar{b} = 12.0$; $\bar{b} = 8.9$). Where the response rate was less than predicted, this was not necessarily due to a failure to meet the ESS minimum requirements regarding contact efforts or indeed due to lack of effort generally (Philippens and Billiet 2003).

5.4. Predictions of Combined DEFF

The prefieldwork predictions of the combined design effect (3) are compared with the predictions that would have been made $(D\hat{E}FF)$ had the realised values of $\{w_i\}$ and \bar{b} been

used in (1) and (2) respectively (Table 2). This is informative of the effect on the sample size calculation of inaccurate prediction of $\{w_i\}$ and \bar{b} . It is not informative of possible differences between the predicted values and values that may be estimated for specific y. Such differences could arise for several reasons, including differences between the predicted value of ρ and the sample-estimated value for any particular y, the level of association between y_{ij} and w_{ij} , and the effect of stratification – all of which could be different for different y. Comparisons between predicted DEFFs and estimated DEFFs for specific y would not in our opinion convey much information unless the differences are decomposed into their component parts. This is a considerable task deserving of separate treatment and best tackled when more than one round of ESS data is available. Sensitivity to the assumptions inherent in (1) and (2) is addressed by Lynn and Gabler (2005).

Differences between the predicted and estimated values of DEFF are nonexistent in some cases, but considerable in others. Three nations are trivial cases as the sample designs involved neither clustering nor unequal selection probabilities (DK, FI, SE). In all other cases there was some uncertainty regarding the parameters of clustering, design weights, or both. In five nations, the uncertainty only concerned \bar{b} . These were three nations (BE, HU, SI) with an equal-probability sample selected from population registers and two (DE, PL) where the weights were completely determined by the sample design. The prediction turned out accurate in Belgium. In Slovenia, \bar{b} was under-estimated as both the eligibility rate and response rate turned out higher than predicted. These two rates were also both under-estimated in Hungary, but this was more than compensated for by an increase in the number of PSUs (and associated reduction in the selected cluster sample

Table 2. Predicted and Estimated Values of the Combined DEFF

	DĔFF	DÊFF	DEFF/DEFF
AT (Austria)	1.38	1.40	0.99
BE (Belgium)	1.10	1.10	1.00
CH (Switzerland)	1.47	1.41	1.04
CZ (Czech Rep)	1.28	1.36	0.94
DE (Germany)	1.52	1.49	1.02
DK (Denmark)	1.00	1.00	1.00
ES (Spain)	1.39	1.42	0.98
FI (Finland)	1.00	1.00	1.00
FR (France)	1.58	1.42	1.11
GR (Greece)	1.36	1.45	0.94
HU (Hungary)	1.19	1.14	1.04
IE (Ireland)	1.24	1.26	0.98
IL (Israel)	1.59	1.92	0.83
IT (Italy)	1.53	1.50	1.02
LU (Luxemburg)	1.40	1.26	1.11
NL (The Netherlands)	1.19	1.19	1.00
NO (Norway)	1.18	1.36	0.87
PL (Poland)	1.12	1.14	0.98
PT (Portugal)	1.19	1.98	0.60
SE (Sweden)	1.00	1.00	1.00
SI (Slovenia)	1.40	1.46	0.96
UK (United Kingdom)	1.55	1.50	1.03

size), subsequent to the prediction made on the sign-off form. Response rate was lower than predicted in Germany, with a consequent over-prediction of *DEFF*, and the opposite was true in Poland.

There were two nations (LU, NL) where the only uncertainty concerned $\{w_i\}$ as the sample was not clustered. In the Netherlands, the design weights depended only on the distribution of household size and this was well known from previous surveys. In Luxemburg, there were two sources of variation in the weights: a deliberate oversampling of certain domains and an unavoidable many-to-many correspondence of frame units to eligible persons within households. The latter could not be well predicted as the relevant information had not been recorded on previous surveys that had used this frame. The prediction of the impact on the design effect turned out to have been pessimistic.

In the remaining nations, there was some uncertainty about both components of DEFF. In three (CZ, ES, FR) $D\tilde{E}FF_P < D\hat{E}FF_P$, but $D\tilde{E}FF_C > D\hat{E}FF_C$ (see Section 5.3. above), so the effect of the latter reduced that of the former. In Israel and Portugal, too, $D\tilde{E}FF_P < D\hat{E}FF_P$, but there was no counter-balancing over-prediction of $DEFF_C$, so overall there was a substantial under-prediction of DEFF, with the consequence that the Israeli and Portuguese samples fail to achieve the ESS effective sample size requirement, even though they had been designed to achieve the requirement and the number of completed interviews exceeded the predicted number. In Norway, too, DEFF was underpredicted, but in this case it was due to under-predicting the effect of clustering $(D\tilde{E}FF_P \cong D\hat{E}FF_P)$, but $D\tilde{E}FF_C < D\hat{E}FF_C)$. In the other nations (AT, CH, FR, GR, IE, IT, UK) the predictions proved reasonably accurate.

Finally, we should note that in four nations (BE, DE, PL, UK) a dual-design was used, involving one stratum where an unclustered sample was selected (the largest cities/municipalities in the case of Belgium, Germany and Poland; Northern Ireland in the case of the United Kingdom) and a second stratum where a clustered sample was selected. The ESS had not provided guidelines on how to predict the combined (national) *DEFF* in such situations and consequently a different method was used in each case. We believe that a weighted combination of two separately predicted design effects should be used in such cases, but the weights should depend on between-strata differences in the overall sampling fraction and in anticipated coverage/response rates (Gabler et al. 2006).

5.5. Predictions of ρ

Only a small number of nations used a value other than $\tilde{\rho}=0.02$ to calculate $D\tilde{E}FF_C$. In these few cases, the values of $\tilde{\rho}$ ranged from 0.03 to 0.05. These values were chosen either because previous "similar" surveys had produced estimates of this general magnitude or because the proposed cluster units were particularly small geographical areas. The ESS Round 1 data now provide the opportunity for estimates of ρ to be compared across countries. This will be done in order to inform the design of future rounds but is not treated here.

5.6. Predictions of Response and Eligibility Rates

As mentioned already in Sections 5.3 and 5.4, predictions of response and eligibility rates were rather inaccurate in some nations. This affected \tilde{b} and hence $D\tilde{E}FF_C$. In no case was

the eligibility rate over-estimated: the tendency was to be over-cautious in cases where the properties of the frame were not well known in relation to the ESS target population. But response rate was over-estimated by 5 percentage points on average and in some cases the over-estimation was much larger (seven countries made a slight under-estimation). Over-estimation persisted even in countries where the predicted response rate was rather low. This may have been caused by the target response rate set (Section 3.3). Participating nations who were sure that they would not meet the target may have been unwilling to admit the anticipated extent of their under-achievement. This is likely to change on future rounds of the survey, now that response rates for all nations for Round 1 have been published and openly discussed in meetings of the NCs. In bilateral discussions, nations that have previously failed to achieve 70% response will be encouraged to aim for modest and realistic improvement.

5.7. Completeness of Documentation

As introduced in Section 3.6, three types of documentation were required:

- 1 Micro-level indicators of selection probabilities, PSU membership and stratum membership;
- 2 A detailed description of the sample design and selection process;
- 3 A description of any ways in which the design falls short of the strict ESS requirements, such as under-coverage of the sampling frame.

The type 1 documentation was supplied accurately and in a form consistent with the specification by most nations. However, in a minority of cases the data supplied initially were found by the sampling panel to be in error. This was due either to a misunderstanding of how the data items should be defined (this applied particularly to conditional selection probabilities) or to a misconception regarding the sample design (the person providing the data was not necessarily the same person with whom the sampling panel had discussed the design). To identify these errors required careful analysis by the sampling panel (akin to "edit checks" typically run on survey data) and to correct them required extended liaison with NCs and survey organisations. Regarding selection probabilities, the specification asked only for the conditional probabilities at each stage, on the assumption that it would reduce the burden on data providers if the sampling panel were to calculate the overall probabilities and hence the design weights. However, for some designs the overall probabilities or the product of those arising from two stages would have been easier to provide and would also have provided a check on the components. With hindsight, asking the data providers to also provide the overall probability would most likely have averted some of the errors and made the process of error detection easier. This will be introduced on future rounds of the ESS.

The type 2 documentation was in most cases a natural by-product of the extended dialogue between the sampling panel member and the NC/survey organisation. The panel member was given the responsibility for producing a final document which was typically a synthesis of information provided and amended over a series of emails and other contacts. This worked well, as it was possible to ensure consistency of content and style while also ensuring that the documentation was produced by someone who had a detailed knowledge

of the design. In a small number of cases, elements of the design changed after the documentation had been produced (which was at the "sign-off" stage). For future rounds, it may be desirable to produce two versions of the document – a preliminary (signed-off) version and a final (post-implementation) version. The former could be less detailed than the latter.

Type 3 documentation proved trivial for some aspects of sample design but a considerable challenge for others. An example of the former is sample size. There were two nations in which the selected sample size was clearly insufficient to provide the required minimum number of interviews. This happened solely because of national budgetary constraints. An example of where the documentation proved to be a challenge is sampling frame coverage. It was rare that a survey organisation or NC had detailed knowledge of the coverage properties of the frame. Rather, they were able to provide only general statements regarding the circumstances that would have led to omission from the frame – but not estimates of the proportion or characteristics of population units omitted. In some cases, sampling panel members were fruitfully able to suggest where such estimates may be obtained, but typically this was not possible and only the general statements could be provided in the documentation. In practice, the documentation is therefore only of limited use in the assessment of likely survey error.

5.8. Examples of Improvement in National Procedures

In some countries, the implemented design represented a significant breakthrough in survey standards. For example, in the Czech Republic a strict probability general population sample was implemented by a private sector survey company for the first time; in Greece an area-based probability sample was used for the first time, and with almost complete coverage (only the smallest islands were excluded), which is unusual; in Italy a new method of including nonelectors in a sample drawn from the electoral registers was developed (Lynn and Pisati 2007); and in France, a probability design was used for the first time in the experience of the survey organisation and researchers involved. In several other nations, more modest advances were made.

In Switzerland, the design incorporated an experiment designed to inform possible improvements in procedures in future years. The ESS specification required a face-to-face initial approach by interviewers to all sample persons/addresses. (Once initial contact had been made, subsequent contacts by telephone were permitted, for example to arrange appointments.) However, common practice in Switzerland is to make the first approach by telephone, for cost-efficiency reasons. As a compromise, a random proportion of the sample addresses were allocated to a "telephone" treatment, while the remainders were approached face-to-face. The intention was that the data resulting from this experiment could be used to assess the extent to which the ESS assumption, that refusal rates would be higher with an initial contact by telephone, held. Preliminary analysis appears in Joye and Bergman (2003). This is an example of the ESS extending methodological knowledge.

6. Conclusion

The aims of the ESS, in terms of sample design standards and procedures for implementation of those standards, were ambitious. Though not realised in every detail,

the ESS can be considered a great success. In particular, the process for developing and finalising sample designs can be considered successful both at a subjective level and in objective terms (guidelines used to estimate design parameters proved useful and estimates generally accurate; documentation is relatively complete).

The evaluation of the estimation of design parameters presented here has provided several pointers to how such estimation might be improved on future cross-national surveys. The nature of uncertainty in the estimates has been described and the directions of errors documented. For example, uncertainty in the distribution of household size was found to be common for address-based samples (Section 5.2). For repeating surveys, such as the ESS, the sample distribution observed at round t might provide a good estimate of the expected distribution at round t+1. For new surveys, analysis of data from other sources (e.g., other surveys or census microdata) might be warranted. Design effect estimates were found to be quite sensitive to predictions of \bar{b} , especially when \bar{b} was expected to be small (Sections 5.3 and 5.6). In this case, inaccuracy in the prediction of eligibility and response rates can produce relatively large shifts in $(\bar{b}-1)$. It can be suggested that conservative estimates of \bar{b} should be used when \bar{b} is small. The importance of guidance for the estimation of design effects with dual designs (Section 5.4) has also been highlighted.

The ESS has provided advances in survey practice in a number of specific nations (Section 5.8). Additionally, we believe that the procedures for sample design that are the subject of this article represent a useful advance in the methodology of cross-national surveys. The documentation of the sample design (Sections 3.6 and 5.7) provides researchers with considerable data for addressing methodological research questions as well as the opportunity to produce variance estimates that appropriately take into account the sample design. This is important for any survey, but particularly difficult to achieve on a cross-national survey. The ESS provides a model for how it can be achieved successfully.

Finally, much of the information presented in this article can, and will, be treated as quality indicators that will feed into a process of continuous quality improvement for the ESS. The sample design procedures have already been amended in the light of the Round 1 experiences and will continue to be reviewed.

7. References

Adams, R. and Wu, M. (eds) (2002). PISA 2000 Technical Report. Paris: Organisation for Economic Co-operation and Development.

Chapman, D.W. (1983). The Impact of Substitution on Survey Estimates. In Incomplete Data in Sample Surveys, Vol.II, Theory and Bibliographies, W.G. Madow, I. Olkin, and D.B. Rubin (eds). New York: Academic Press.

Cornelius, R. (1985). The World Fertility Survey and Its Implications for Future Surveys. Journal of Official Statistics, 1, 427–433.

Elliot, D. (1993). The Use of Substitution in Sampling. Survey Methodology Bulletin, 33, 8–11.

European Science Foundation (1999). The European Social Survey (ESS) – A Research Instrument for the Social Sciences in Europe, Strasbourg: European Science Foundation. Gabler, S., Häder, S., and Lahiri, P. (1999). A Model Based Justification of Kish's Formula for Design Effects for Weighting and Clustering. Survey Methodology, 25, 105–106.

- Gabler, S., Häder, S., and Lynn, P. (2006). Design Effects for Multiple Design Samples. Survey Methodology, 32, 115–120.
- Häder, S. and Gabler, S. (2003). Sampling and Estimation. In Cross Cultural Survey Methods, J. Harkness, F. van de Vijver, and P. Mohler (eds). New York: John Wiley and Sons.
- Joye, D. and Bergman, M. (2003). ESS in Switzerland and Modes of Recruitment: Some Remarks. Paper presented to the Methodology Committee of the European Social Survey, London, 14 November.
- Kish, L. (1987). Weighting in Deft². The Survey Statistician, June.
- Kish, L. (1992). Weighting for Unequal P_i. Journal of Official Statistics, 8, 183–200.
- Kish, L. (1994). Multipopulation Survey Designs: Five Types with Seven Shared Aspects. International Statistical Review, 62, 167–186.
- Kish, L. (1995). Methods for Design Effects. Journal of Official Statistics, 11, 55-77.
- Le, T. (1993). Sampling Practice in the Demographic and Health Surveys. Bulletin of the International Statistical Institute, Contributed Papers, Book 2, 103–104.
- Le, T. and Verma, V. (1997). An Analysis of Sampling Designs and Sampling Errors of the Demographic and Health Surveys. DHS Analytic Report No.3, Calverton, MD: Macro International Ltd. http://www.measuredhs.com/pubs/details.cfm?ID = 4
- Lyberg, L. (2000). Review of IALS A Commentary on the Technical Report. In Measuring Adult Literacy. The International Adult Literacy Survey in the European Context. London: Office for National Statistics.
- Lynn, P. (2003). Developing Quality Standards for Cross-national Survey Research: Five Approaches. International Journal of Social Research Methodology, 6, 323–336.
- Lynn, P. (2004). The Use of Substitution in Surveys. The Survey Statistician, 49, 14–16. Lynn, P. and Gabler, S. (2005). Approximations to *b** in the Prediction of Design Effects due to Clustering. Survey Methodology, 31, 101–104.
- Lynn, P. and Pisati, M. (2007). Improving the Quality of Sample Design for Social Surveys in Italy. Journal of the Italian Statistical Society. (Forthcoming)
- Lynn, P., Japec, L., and Lyberg, L. (2006). What's So Special About Cross-national Surveys? In CSDI 2005: Papers from the Third International Workshop on Comparative Survey Design and Implementation, J.A. Harkness (ed.). ZUMA Nachrichten Spezial 12, Mannheim: ZUMA.
- Philippens, M. and Billiet, J. (2003). Nonresponse and Fieldwork Efforts in the ESS: Results from the Analysis of Call Record Data. Paper presented to the Methodology Committee of the European Social Survey, London, 14 November.
- Rubin, D.B. and Zanutto, E. (2002). Using Matched Substitutes to Adjust for Nonignorable Nonresponse Through Multiple Imputations. In Survey Nonresponse, R.M. Groves, D.A. Dillman, J.L. Eltinge, and R.J.A. Little (eds). New York: John Wiley and Sons.
- Särndal, C.-E., Swensson, B., and Wretman, J. (1992). Model-assisted Survey Sampling. New York: Springer-Verlag.
- Vehovar, V. (2003). Field Substitutions Redefined. The Survey Statistician, 48, 35–37.